# An Unbiased Explicit Adaptive Gain and Time-delay Estimation Algorithm

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#### • Motivation

- Adaptive Gain-Delay Estimation
- Unbiased Algorithm
- Simulations
- Conclusion



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#### Motivation

- Some applications of gain and time-delay adaptive estimation, such as 2-channel time-interleaved ADC calibration require low complexity and high precision estimates.
- Low complexity  $\Rightarrow$  need to use an **explicit algorithm**.
- A common explict gain and time-delay adaptive estimation algorithm, due to Kong and Solo (1991), exhibits a gain bias inversely proportional to SNR.
- Source of the bias: the power of the noise component of the error is **gain dependent**.





- By **normalizing the error** an unbiased adaptive explicit gain and time-delay algorithm is proposed.
- Unbiasedness is demonstrated both theoretically by using the ODE method to analyze convergence of the algorithm, and through simulations.





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#### **Adaptive Gain-Delay Estimation**

Consider observations

$$y_1(n) = s(n) + v_1(n)$$
  
 $y_2(n) = h(n, \theta) * s(n) + v_2(n) ,$ 

where s(n) = unknown signal,  $v_i(n) =$  independent zero-mean WGNs with variance v for i = 1, 2,

$$\boldsymbol{\theta} = \begin{bmatrix} g \\ d \end{bmatrix} =$$
 unknown parameter vector

with g = gain d = delay, and

$$h(n, \boldsymbol{\theta}) = g \frac{\sin(\pi(n-d))}{\pi(n-d)}$$

= impulse response of gain-delay filter  $H(e^{j\omega}, \theta) = ge^{-j\omega d}$ .

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#### Consider error

$$e_{\mathrm{B}}(n, \hat{\theta}) = y_2(n) - h(n, \hat{\theta}) * y_1(n)$$
  
=  $y_2(n) - \hat{g}y_1(n - \hat{d}),$ 

where

$$\hat{\boldsymbol{\theta}} = \begin{bmatrix} \hat{g} \\ \hat{d} \end{bmatrix} = \text{estimated parameter vector}$$

The explicit adaptive gain-delay algorithm proposed by Kong and Solo (1991) uses a stochastic gradient algorithm to minimize

$$J_{\rm B}(\hat{\boldsymbol{\theta}}) = \frac{1}{2} E[e_{\rm B}^2(\hat{\boldsymbol{\theta}})]$$

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This gives

$$\hat{\boldsymbol{\theta}}(n+1) = \hat{\boldsymbol{\theta}}(n) + \mu_{\mathrm{B}}e_{\mathrm{B}}(n,\hat{\boldsymbol{\theta}})\mathbf{q}(n,\hat{\boldsymbol{\theta}}),$$

where  $\mu_{\rm B} =$  step size, and

$$\mathbf{q}(n,\hat{\boldsymbol{\theta}}) = \begin{bmatrix} q_1(n,\hat{\boldsymbol{\theta}}) \\ q_2(n,\hat{\boldsymbol{\theta}}) \end{bmatrix} = -\nabla_{\hat{\boldsymbol{\theta}}} e_{\mathrm{B}}(n,\hat{\boldsymbol{\theta}})$$
$$= \nabla_{\hat{\boldsymbol{\theta}}} h(n,\hat{\boldsymbol{\theta}}) * y_1(n)$$

with

$$\nabla_{\hat{\theta}} \stackrel{\triangle}{=} \begin{bmatrix} \frac{\partial}{\partial \hat{g}} & \frac{\partial}{\partial \hat{d}} \end{bmatrix}^T$$
  
= gradient with respect to vector  $\hat{\theta}$ .

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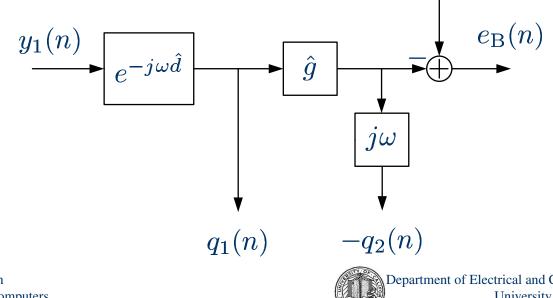
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In the frequency domain

$$\nabla_{\hat{\theta}} H(e^{j\omega}, \hat{\theta}) = \begin{bmatrix} 1\\ -j\omega \hat{g} \end{bmatrix} e^{-j\omega \hat{d}},$$

so that  $e_B(n)$ ,  $q_1(n)$  and  $q_2(n)$  can be obtained as shown below.  $y_2(n)$ 



**Bias evaluation:** Assume unknown signal s(n) is WSS with PSD  $S(e^{j\omega})$  and total power P. The error

$$e_{\rm B}(n, \hat{\theta}) = gs(n-d) - \hat{g}s(n-\hat{d}) + v_2(n) - \hat{g}v_1(n-\hat{d}),$$

SO

$$J_{\rm B}(\hat{\theta}) = \frac{1}{4\pi} \int_{-\pi}^{\pi} |ge^{-j\omega d} - \hat{g}e^{-j\omega \hat{d}}|^2 S(e^{j\omega}) d\omega$$
$$+ (1 + \hat{g}^2)\frac{v}{2} .$$

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In the vicinity of g and d

$$J_{\rm B}(\hat{\theta}) \approx \frac{A}{2}(\hat{g} - g)^2 + \frac{B}{2}(\hat{d} - d)^2 + (1 + \hat{g}^2)\frac{v}{2} ,$$

with

$$A = \frac{1}{2} \int_{-\pi}^{\pi} S(e^{j\omega}) d\omega = P$$
$$B = \frac{1}{2} \int_{-\pi}^{\pi} \omega^2 S(e^{j\omega}) d\omega .$$

If SNR = P/v, it is minimized by

$$\hat{g} = \frac{g}{1 + (\text{SNR})^{-1}} \approx [1 - (\text{SNR})^{-1}]g$$

and  $\hat{d} = d$ . The gain estimate is biased!

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### **Unbiased Algorithm**

To remove the estimated gain dependence of the noise component of the error, use the scaled error

$$e(n, \hat{\theta}) = \frac{1}{(1 + \hat{g}^2)^{1/2}} e_{\mathrm{B}}(n, \hat{\theta})$$
  
=  $\frac{1}{(1 + \hat{g}^2)^{1/2}} [y_2(n) - \hat{g}y_1(n - \hat{d})],$ 

and use a stochastic gradient scheme to minimize

$$J(\hat{\boldsymbol{\theta}}) = \frac{1}{2} E[e^2(n, \hat{\boldsymbol{\theta}})] = \frac{1}{1 + \hat{g}^2} J_{\mathrm{B}}(\hat{\boldsymbol{\theta}})$$

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# **Unbiased Algorithm (cont.)**

- The scale factor  $(1 + \hat{g}^2)^{-1}$  distorts the shape of objective function  $J(\hat{\theta})$ and may change the domain where it is convex.
- In the vicinity of  $\boldsymbol{\theta}, J(\hat{\boldsymbol{\theta}})$  admits the approximation

$$J(\hat{\theta}) \approx \frac{1}{2} \Big( \frac{1}{(1+g^2)} [A(\hat{g}-g)^2 + B(\hat{d}-d)^2] + v \Big) ,$$

whose minimum is  $\theta \Rightarrow$  new algorithm will be unbiased.



#### **Unbiased Algorithm (cont.)**

#### **Unbiased adaptive algorithm:**

$$\hat{\boldsymbol{\theta}}(n+1) = \hat{\boldsymbol{\theta}}(n) - \mu e(n, \hat{\boldsymbol{\theta}}) \nabla_{\hat{\boldsymbol{\theta}}} e(n, \hat{\boldsymbol{\theta}}) ,$$

where  $\mu = \text{step size}$ , and

$$\frac{\partial e}{\partial \hat{g}} = -\frac{1}{(1+\hat{g}^2)^{3/2}} [\hat{g}y_2(n) + y_1(n-\hat{d})] \frac{\partial e}{\partial \hat{d}} = -\frac{1}{(1+\hat{g}^2)^{1/2}} \frac{\partial h}{\partial \hat{d}}(n,\hat{\theta}) * y_1(n) .$$

For small  $\hat{d}$ , the time-varying delay filter  $e^{-j\omega\hat{d}}$  can be implemented with Farrow's expansion

$$e^{-j\omega\hat{d}} \approx 1 - j\omega\hat{d} + (-j\omega)^2 \frac{(\hat{d})^2}{2!} + \ldots + (-j\omega)^k \frac{(\hat{d})^k}{k!},$$

where k = approximation order.

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# **Unbiased Algorithm (cont.)**

**Convergence analysis:** By using the ODE/ stochastic averaging method, the sample paths  $\hat{\theta}(n)$  of the parameter estimates are approximated by the solutions of ODE

$$\frac{d\hat{\boldsymbol{ heta}}}{dt}(t) = -\nabla_{\hat{\boldsymbol{ heta}}} J(\hat{\boldsymbol{ heta}}(t)) \; .$$

Starting from an initial value  $\hat{\theta}(0)$  sufficiently close to  $\theta$ , all trajectories converge to  $\theta$ , and for small  $\mu$  and large n, the estimated parameter vector  $\hat{\boldsymbol{\theta}}(n)$  is  $N(\boldsymbol{\theta}, \mu \mathbf{P})$  distributed.

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#### Simulations

- Assume s(n) formed by the supperposition of 3 unit power sinusoids with frequencies 0.1π, 0.4π, and 0.8π and uniformly distributed random phases.
- Noise variance  $v = 6 \times 10^{-2}$ , so SNR = 50 (17dB).
- True gain g = 0.9 and delay d = 0.1 (10% mismatches between the two channels).
- Farrow expansion of order 2 used for delay filter. Differentiator  $j\omega$ implemented by applying Kaiser window of length 81 and  $\beta = 4$  to ideal impulse response

$$k(n) = \begin{cases} 0 & n = 0\\ \frac{(-1)^n}{n} & n \neq 0 \end{cases}.$$

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## **Simulations (cont.)**

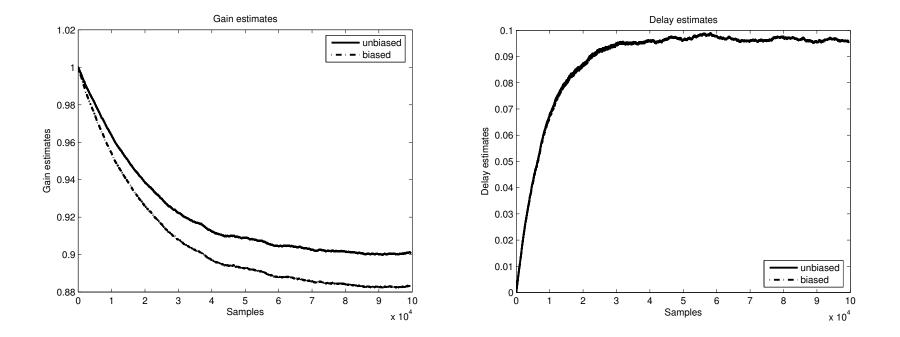
- Compare biased and unbiased algorithms with  $\mu_{\rm B} = 0.5\mu$  and  $\mu = 10^{-4.5}$  (factor 1/2 Due to the relative scaling of  $J_{\rm B}$  and J).
- Use  $\hat{g}(0) = 1$  and  $\hat{d}(0) = 0$  as initial estimates.
- Computed biased gain estimate:  $\hat{g}_{\rm B} = 0.8851$ , close to the value

$$\frac{1}{1 + (\mathrm{SNR})^{-1}}g = 0.8824$$

predicted by theoretical analysis.



#### **Simulations (cont'd)**



Estimated gains and delays for 3 sinusoids signal and 17dB SNR





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### Conclusion

- By applying a  $(1 + \hat{g}^2)^{-1}$  scaling to the objective function used by the Kong-Solo explicit gain-delay estimation algorithm, an unbiased algorithm was derived.
- The scaling shifts the local minimum of objective function J(θ̂) to the true value θ, but distorts the function far away from θ.
- The algorithm is best adapted to cases (like TIADC calibration) where an initial estimate  $\hat{\theta}(0)$  close to  $\theta$  is available.





#### Thank you!!

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